

Causal Discovery and Inference in Time Series

C. K. Assaad, E. Devijver, *E. Gaussier*, G. Goessler, A. Meynaoui, L. Zan

17 April 2023

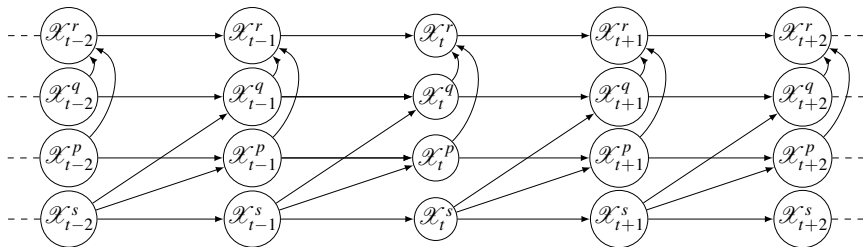
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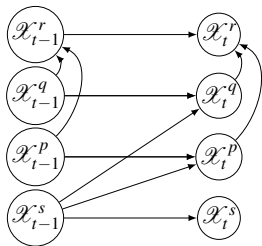
Time series are everywhere

- Time series arise as soon as observations, from sensors or experiments, for example, are collected over time
- They are present in various forms in many different domains
 - Healthcare (through, e.g., monitoring systems)
 - Industry 4.0 (through, e.g., predictive maintenance and industrial monitoring systems)
 - Surveillance systems (from images, acoustic signals, seismic waves, etc.)
 - Energy management (through, e.g. energy consumption data)
 - ...

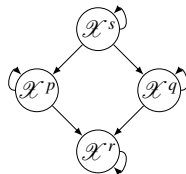
How to represent relations in time series?



Full Time Causal Graph (a)



Window Causal Graph (b)



Summary Causal Graph (c)

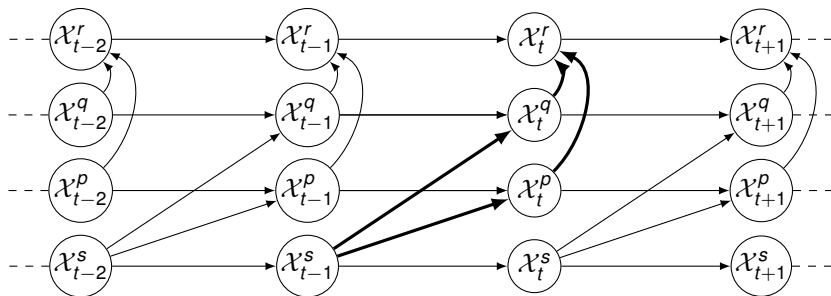
Causal graphs in time series (1)

Assumption 1 - Temporal priority *A cause does not occur after its effects (maximal temporal lag)*

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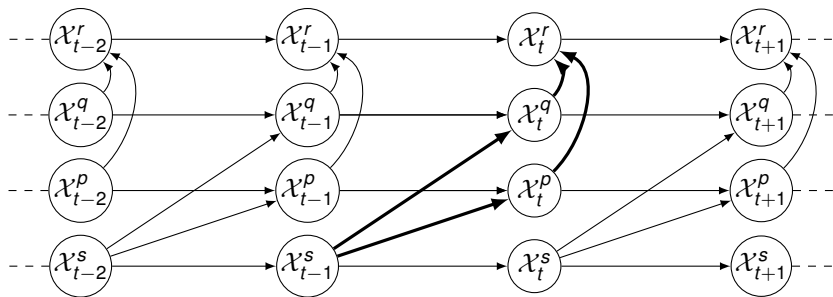
The full story



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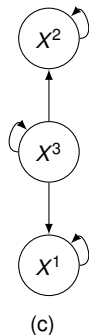
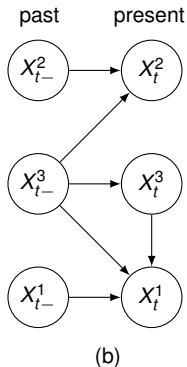
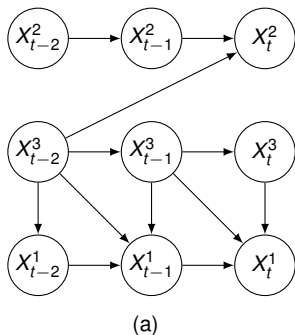
The full story



Assumption 2 - Consistency through time [causal stationarity] *All causal relationships remain constant in direction throughout time*

Causal graphs in time series (2)

Summarizing the full-time window graph with or without loss of information: window causal graph (a), extended summary causal graph (b) and summary causal graph (c)



Causal graphs in time series (3)

Remarks

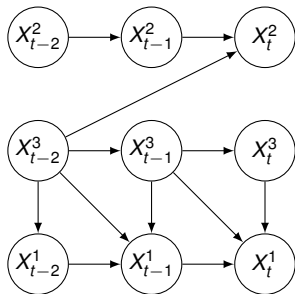
- Window causal graphs equivalent to full-time graphs (cons. through time)
- Unique extended summary causal graph for a given window causal graph (reverse not true)
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- Advantage/disadvantage of summary causal graphs
 - + Easier to manipulate by experts
 - Less precise than window causal graphs

Causal graphs in time series (3)

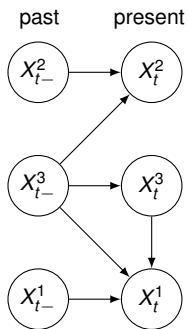
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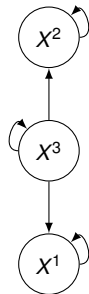
Assumption 3 - Acyclicity *All causal graphs are acyclic*



(a)



(b)



(c)

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Focus on summary causal graphs (starting with extended summary causal graphs) and continuous time series

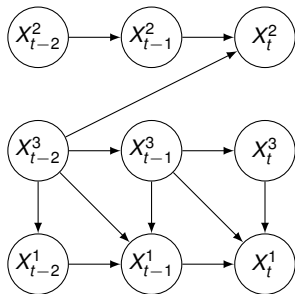
Problem statement, Markov cond. and faithfulness

Problem: *directly* infer extended summary graph from observational data (causal structure learning)

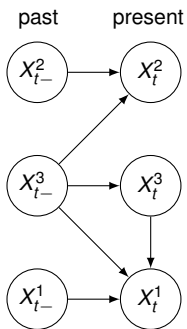
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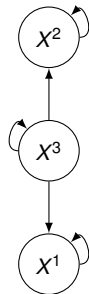
Markov condition *A necessary and sufficient condition for a probability distribution to be compatible with a DAG \mathcal{G} is that every variable be independent of all its nondescendants (in \mathcal{G}), conditional on its parents*



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(b)



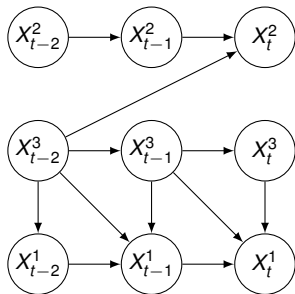
(c)

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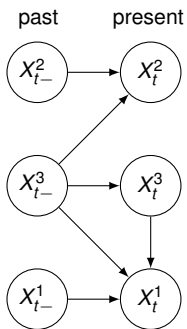
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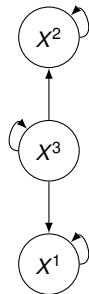
Assumption 4 - Faithfulness *We say that a graph \mathcal{G} and a compatible probability distribution P are faithful to one another if all and only the conditional independence relations true in P are entailed by the Markov condition applied to \mathcal{G}*



(a)



(b)



(c)

Truly causal methods?

General functional model of any potential effect X^q

$$X_t^q = f(C_t^q(X^{r_1}), \dots, C_t^q(X^{r_q}), \xi_t^q)$$

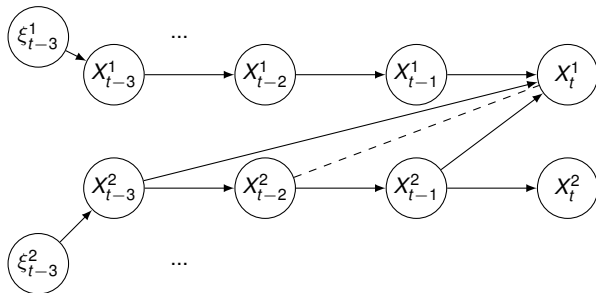
- f : any real-valued multivariate function
- C^q : set of causes, $C_t^q(X^r)$: past instants of X^r which are actual causes
- ξ_t^q : noise independent from all causes

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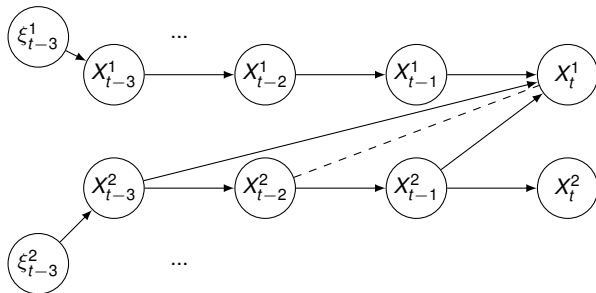


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Truly causal methods aim at distinguishing causal correlations from spurious correlations

Families of causal methods

Main families of causal methods

- Granger causality, CCM causality, PAI causality
- Constraint-based approaches, noise-based approaches, score-based approaches, logic-based approaches

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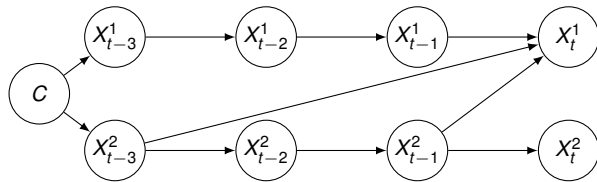
Focus on constraint-based approaches (and noise-based approaches):
popular, sound and complete, no assumption on f

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PC with causal sufficiency

Assumption 5 - Causal sufficiency *Observed variables are causally sufficient, i.e., all common causes of all variables are observed*



PC with causal sufficiency

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The PC algorithm (Spirtes-2000; Colombo-2014)

- 1 Build complete undirected graph U between all pairs of variables
- 2 Set $n = 0$
- 3 For each pair of adjacent variables (X, Y) in (current) U such that $ADJ(U, X) - Y$ or $ADJ(U, Y) - X$ has at least n elements, check conditional independence bet. X and Y with any subset of S of n elements. If $X \perp\!\!\!\perp Y | S$, $S = Sepset(X, Y)$
- 4 $n = n + 1$
- 5 Execute orientation rules on obtained skeleton using temporal priority between the past and present slices

Conditional independence at the core of the procedure for constructing the skeleton

Independence measure: mutual information

Mutual information ...

- ... has proven useful in several studies on causal inference (Affeldt-2015, Runge-2019, Runge-2020)
- ... does not require assumptions on the generative model
- ... can be (somewhat) robustly estimated from observational data (k -NN)

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Present slice Standard mutual information

Between past and present slices X^p does not directly cause X^q if there exists a subset of time series \mathcal{X}^R such that $\forall K \in \mathbb{Z}^*$,

$\forall \{\gamma_1, \dots, \gamma_K\}$ s.t. $0 \leq \gamma_K < \dots < \gamma_1$:

$$I(X_t^q; X_{t-\gamma_1}^p, \dots, X_{t-\gamma_K}^p | \mathcal{X}_{t^*}^R) = 0$$

Efficient estimation

One can efficiently identify independence between considering the window in X^p starting at $t - \gamma$ and ending at t , denoted $t - \gamma : t$ where γ is the maximal gap:

Efficient estimation

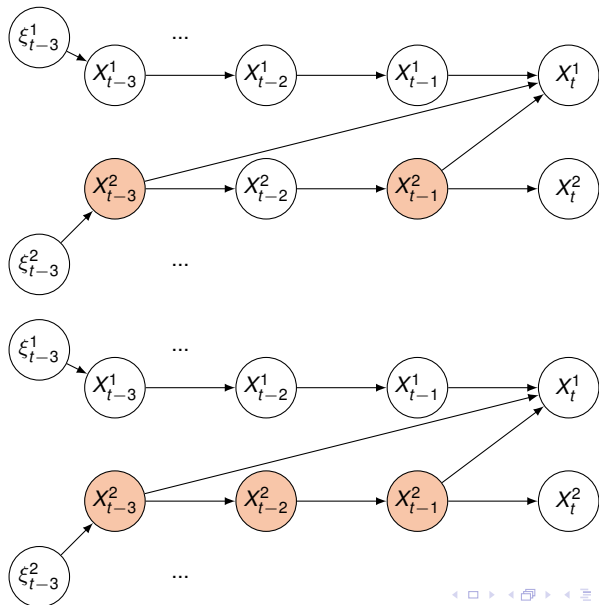
One can efficiently identify independence between considering the window in X^P starting at $t - \gamma$ and ending at t , denoted $t - \gamma : t$ where γ is the maximal gap:

Property

Let γ denote the maximum gap between a cause and its effect. The following two propositions are equivalent:

- (a) $I(X_t^q; X_{t-\gamma_1}^p, \dots, X_{t-\gamma_K}^p | \mathcal{X}_{t^*}^R) = 0, \forall K \geq 1, \forall \gamma_1 > \dots > \gamma_K \geq 0,$
- (b) $I(X_t^q; X_{t-\gamma:t}^p | \mathcal{X}_{t^*}^R) = 0$

Illustration



Greedy causation entropy

Definition (causal greedy entropy)

The greedy causation entropy, denoted by GCE , from the time series X^p to the time series X^q is defined by:

$$GCE(X^p \rightarrow X^q) = I(X_t^q; X_{t-\gamma:t-1}^p) \quad (1)$$

Denoting by X^R a set of m time series $\{X^{r_1}, \dots, X^{r_m}\}$, the conditional greedy causation entropy takes the form:

$$GCE(X^p \rightarrow X^q | X^R) = I(X_t^q; X_{t-\gamma:t-1}^p | X_{t^*}^{r_1}, \dots, X_{t^*}^{r_m}), \quad (2)$$

where t^* denotes either the present instant t or the time window $t-\gamma:t-1$

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Estimation k -NN method (Frenzel-2007)

From a skeleton to a causal graph: orientation rules

PC-Rule 0 - Origin of causality

- (i) In an unshielded triple $X_t^p - X_t^r - X_t^q$, if $X_t^r \notin \text{Sepset}(p \leftrightarrow q)$, then X_t^r is an unshielded collider: $X_t^p \rightarrow X_t^r \leftarrow X_t^q$
- (ii) In an unshielded triple $X_{t-}^q \rightarrow X_t^q - X_t^p$, if $X_t^q \notin \text{Sepset}(q \rightarrow p)$, then X_t^q is an unshielded collider: $X_{t-}^q \rightarrow X_t^q \leftarrow X_t^p$

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PC-Rule 1- Propagation of causality In an unshielded triple $X_t^p \rightarrow X_t^r - X_t^q$ (resp. $X_{t-}^p \rightarrow X_t^r - X_t^q$), if $X_t^r \in \text{Sepset}(p \leftrightarrow q)$ then orient the unshielded triple as $X_t^p \rightarrow X_t^r \rightarrow X_t^q$ (resp. $X_{t-}^p \rightarrow X_t^r \rightarrow X_t^q$)

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PC-Rule 2 If there exist a direct path from X_t^p to X_t^q and an edge between X_t^p and X_t^q , then orient $X_t^p \rightarrow X_t^q$

From a skeleton to a causal graph: orientation rules

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PC-Rule 3 Orient $X_t^p - X_t^q$ as $X_t^p \rightarrow X_t^q$ whenever there are two paths $X_t^p - X_t^r \rightarrow X_t^q$ and $X_t^p - X_t^s \rightarrow X_t^q$

Wrapping-up

- Simple algorithm with causal sufficiency:
 - Standard skeleton construction based on greedy causation entropy
 - More or less standard orientation rules operating on the present slice
- Theoretical guarantees under assumptions made

Theorem

Let the distribution of V be faithful to a DAG $\mathcal{G} = (V, E)$, and assume that we are given perfect conditional independence information about all pairs of variables (X^p, X^q) in V given subsets $X^R \subseteq V \setminus \{X^p, X^q\}$. Then the skeleton constructed previously followed by the above orientation rules represents the CPDAG of the extended summary causal graph \mathcal{G}

FCI without causal sufficiency

- 1 Same skeleton construction as before
- 2 Orientation rules
 - FCI-Rule 0 (origin of causality) is adapted as before
 - FCI-Rules 1, 2, 3 and 4 (Spirtes-2000)
 - FCI-Rules 8, 9 and 10 (Zhang-2008) - rules 5, 6 and 7 deal with selection bias

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Remark Possible D-separation sets vs. separation sets, PAG (MAG) vs CPDAG (DAG)

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Experimental setting

Datasets, evaluation measures, methods

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- Simple artificial datasets correspond to three different causally sufficient structures and two non causally sufficient structures generated by:

$$X_t^q = a_{t-1}^{qq} X_{t-1}^q + \sum_p a_{t-\gamma}^{pq} f(X_{t-\gamma}^p) + 0.1 \xi_t^q$$

with $\gamma \geq 0$, $a_t^{jq} \sim \mathcal{U}([-1; -0.1] \cup [0.1; 1])$, $\xi_t^q \sim \mathcal{N}(0, 1)$ and f chosen at random in {absolute value, tanh, sine, cosine}

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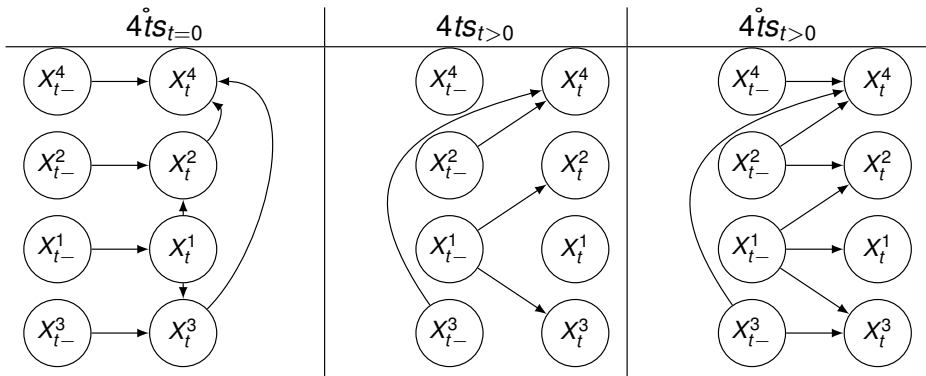
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- FMRI datasets for 28 different brain networks (Smith-2011)
- F1-score (between and within time series)
- PCMICI (Runge-2019, Runge-2020), oCSE (Sun-2015), tsFCI (Entner-2010), VarLINGAM (Hyvärinen-2010), Dynotears (Pamfil-2020), GCMVL (Arnold-2017), TCDF (Nauta-2019)

Artificial datasets (extract)



Results

No hidden common cause

	Perf.	PCGCE	oCSE	PCMCI	VarLiNGAM	Dynotears	TCDF	MVGCL
$4\hat{t}s_{t=0}$	$F^{p \neq q}$	0.62 \pm 0.17	–	0.60 \pm 0.12	0.32 \pm 0.13	0.04 \pm 0.12	0.00 \pm 0.00	–
	$F^{p=q}$	0.81 \pm 0.12	–	0.87 \pm 0.12	0.92 \pm 0.07	0.37 \pm 0.21	0.18 \pm 0.24	–
$4\hat{t}s_{t>0}$	$F^{p \neq q}$	0.71 \pm 0.13	0.31 \pm 0.21	0.67 \pm 0.16	0.00 \pm 0.00	0.16 \pm 0.19	0.00 \pm 0.00	0.52 \pm 0.11
	$F^{p=q}$	0.81 \pm 0.18	0.78 \pm 0.17	0.81 \pm 0.12	0.00 \pm 0.00	0.16 \pm 0.19	0.04 \pm 0.12	0.53 \pm 0.09
$4\hat{t}s_{t>0}$	$F^{p \neq q}$	0.94 \pm 0.06	0.82 \pm 0.11	0.97 \pm 0.05	0.98 \pm 0.04	0.47 \pm 0.15	0.35 \pm 0.27	–
	$F^{p=q}$							

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$4\hat{t}s_{t=0}$	$F^{p \neq q}$	0.62 \pm 0.17	–	0.60 \pm 0.12	0.32 \pm 0.13	0.04 \pm 0.12	0.00 \pm 0.00	–
	$F^{p=q}$	0.81 \pm 0.12	–	0.87 \pm 0.12	0.92 \pm 0.07	0.37 \pm 0.21	0.18 \pm 0.24	–
$4\hat{t}s_{t>0}$	$F^{p \neq q}$	0.71 \pm 0.13	0.31 \pm 0.21	0.67 \pm 0.16	0.00 \pm 0.00	0.16 \pm 0.19	0.00 \pm 0.00	0.52 \pm 0.11
	$F^{p=q}$	0.81 \pm 0.18	0.78 \pm 0.17	0.81 \pm 0.12	0.00 \pm 0.00	0.16 \pm 0.19	0.04 \pm 0.12	0.53 \pm 0.09
		0.94 \pm 0.06	0.82 \pm 0.11	0.97 \pm 0.05	0.98 \pm 0.04	0.47 \pm 0.15	0.35 \pm 0.27	–

No hidden common cause (FMRI)

	Perf.	PCGCE	oCSE	PCMCI	VarLiNGAM	Dynotears	TCDF	MVGCL
FMRI	$F^{p \neq q}$	0.31 \pm 0.2	0.16 \pm 0.19	0.22 \pm 0.18	0.49 \pm 0.28	0.34 \pm 0.13	0.06 \pm 0.12	0.35 \pm 0.08

Results

No hidden common cause

	Perf.	PCGCE	oCSE	PCMCI	VarLiNGAM	Dynotears	TCDF	MVGCL
$4\overset{\circ}{t}s_{t=0}$	$F^{p \neq q}$	0.62 \pm 0.17	–	0.60 \pm 0.12	0.32 \pm 0.13	0.04 \pm 0.12	0.00 \pm 0.00	–
	$F^{p=q}$	0.81 \pm 0.12	–	0.87 \pm 0.12	0.92 \pm 0.07	0.37 \pm 0.21	0.18 \pm 0.24	–
$4\overset{\circ}{t}s_{t>0}$	$F^{p \neq q}$	0.71 \pm 0.13	0.31 \pm 0.21	0.67 \pm 0.16	0.00 \pm 0.00	0.16 \pm 0.19	0.00 \pm 0.00	0.52 \pm 0.11
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With hidden common causes

	Perf.	FCIGCE	tsFCI	TCDF
$7\overset{\circ}{t}2h_{t>0}$	$F^{p \neq q}$	0.57 \pm 0.1	0.52 \pm 0.1	0.02 \pm 0.1
$7\overset{\circ}{t}2h_{t>0}$	$F^{p \neq q}$	0.33 \pm 0.1	0.36 \pm 0.1	0.07 \pm 0.1
	$F^{p=q}$	0.83 \pm 0.1	0.99 \pm 0.1	0.19 \pm 0.2

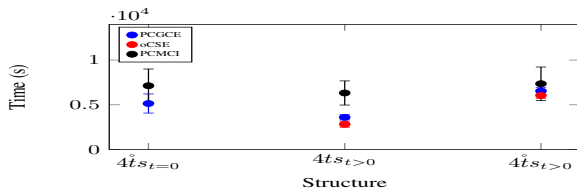
First conclusions

- 1 Relatively simple extensions of PC and FCI to extended summary graphs (Assaad-2022(b))
- 2 Method sound and complete with robust behaviour on various datasets
- 3 Assumptions: temporal priority, consistency through time, acyclicity, faithfulness
- 4 Remark: possibility to mix approaches, *e.g.* noise-based and constraint-based (Assaad-2021)

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Time complexity



What about summary graphs?

The situation in summary graphs is slightly more difficult as there is no distinction between past and present so that windows need be considered on each time series. One can nevertheless define a mutual information measure based on these windows and follow an approach similar to the one above with additional rules for instantaneous relations (Assaad-2022(c)).

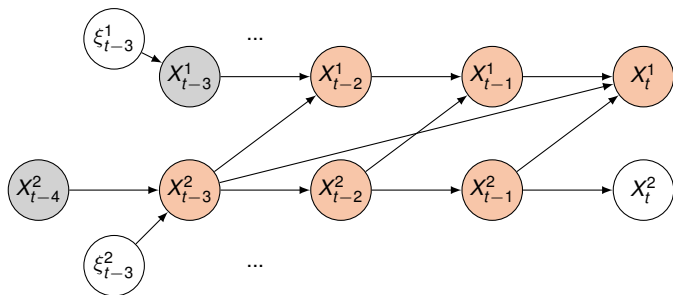


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Background

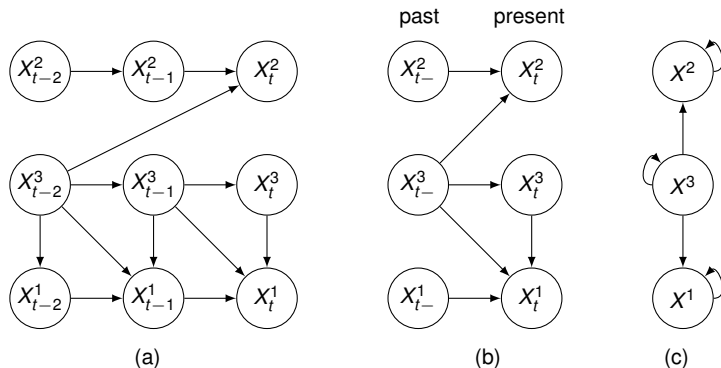
General form of an intervention $P(\mathbf{R}_t = \mathbf{r}_t | do(\mathbf{I}_t = \mathbf{i}_t))$ where \mathbf{I}_t is the intervention set, \mathbf{R}_t the response set and $do()$ the standard intervention operator (Pearl-2009)

Background

$P(X_t^1 = x_t^1, X_t^3 = x_t^3 | do(X_{t-3}^2 = x_{t-3}^2))$ is an example of an intervention where the response set is $\{X_t^1, X_t^3\}$ and the intervention set is $\{X_{t-3}^2\}$

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Problem statement

Identifiability If it exists, provide an equivalent do-free formula, or estimand, for an intervention which can be estimated from observational data - *using observational data to compute a causal effect*

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Example In the above window causal graph,

$P(x_t^1 | do(x_{t-1}^3)) = \sum_{x_{t-2}^3} P(x_{t-2}^3) P(x_t^1 | x_{t-1}^3, x_{t-2}^3)$ as X_{t-2}^3 satisfies the back-door criterion

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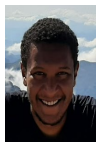
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→ *Anouar's pres. tomorrow*



Equivalence relation

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A simple approach to identifiability is to enumerate all window causal graphs compatible with a given (extended) summary causal graph and use a version of the ID algorithm (Shpitser-2006). However not possible in practice when the number of time series considered is large

Equivalence relation

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Equivalent window causal graph *Let G_s be a (extended) summary causal graph. Consider an intervention set \mathbf{I}_t and a response set \mathbf{R}_t . Two window causal graphs G_{w_1} and G_{w_2} are equivalent for the intervention $P(\mathbf{r}_t | do(\mathbf{i}_t))$ if and only if the intervention is non-identifiable in both graphs, and there is no estimand, or identifiable in both with the same estimand*

Possible approach

- 1 Identify relevant interventional parents blocking all backdoor paths for a given intervention
- 2 Compute do-free formula from relevant interventional parents

Preliminary experimental results encouraging with possibility to use maximal equivalence class in all cases for extended summary graphs

Assumptions & data types

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- 1 Temporal priority, acyclicity, faithfulness
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Data types From purely continuous to **mixed (quantitative and qualitative) data**

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Dealing with mixed data

Once one has a way to estimate the (conditional) mutual information between mixed data samples, then one directly apply the above approach to infer causal graphs from observational data

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Different approaches have been proposed in the literature to estimate the (conditional) entropy of a set of variables, the two prominent ones being based on histograms, particularly adapted to qualitative data, and the k-NN approach, mostly used on quantitative data (Zan-2022)

We propose here an hybrid approach combining both histograms and k-NNs

Decomposing mutual information

Let us consider three mixed random vectors X , Y and Z , where any of their components can be either qualitative or quantitative. Let us denote by X^t (respectively Y^t , Z^t) the sub-vector of X (respectively Y , Z) composed by the quantitative components. Similarly, we denote by X^ℓ (respectively Y^ℓ , Z^ℓ) the sub-vector of qualitative components of X (respectively Y , Z). Then one has:

$$\begin{aligned}
 I(X; Y|Z) &= H(X, Z) + H(Y, Z) - H(X, Y, Z) - H(Z) \\
 &= H(X^t, X^\ell, Z^t, Z^\ell) + H(Y^t, Y^\ell, Z^t, Z^\ell) \\
 &\quad - H(X^t, X^\ell, Y^t, Y^\ell, Z^t, Z^\ell) - H(Z^t, Z^\ell) \\
 &= H(X^t, Z^t|X^\ell, Z^\ell) + H(Y^t, Z^t|Y^\ell, Z^\ell) - H(X^t, Y^t, Z^t|X^\ell, Y^\ell, Z^\ell) \\
 &\quad - H(Z^t|Z^\ell) + H(X^\ell, Z^\ell) + H(Y^\ell, Z^\ell) - H(X^\ell, Y^\ell, Z^\ell) - H(Z^\ell)
 \end{aligned}$$

Estimating $H(., .)$ and $H(.|..)$

$H(., .)$ is estimated with histograms

$H(.^t|.^\ell)$ is estimated with k-NNs

Theorem

Let (X, Y, Z) be a qualitative-quantitative mixed random vector. The obtained estimator $\hat{I}(X; Y|Z)$ is consistent, meaning that, for all $\varepsilon > 0$

$$\lim_{n \rightarrow \infty} P(|\hat{I}(X; Y|Z) - I(X; Y|Z)| > \varepsilon) = 0$$

In addition, $\hat{I}(X; Y|Z)$ is asymptotically unbiased, that is

$$\lim_{n \rightarrow \infty} \mathbb{E}[\hat{I}(X; Y|Z) - I(X; Y|Z)] = 0$$

Illustration (1)

Comparison for $I(X; Y)$ and $I(X; Y|Z)$ (Z quantitative) with one state-of-the-art histogram-based, namely LH (Marx-2021), and three state-of-the-art methods based on k-NN, namely FP (Frenzel-2007), MS (Mesner-2020), RAVK (Rahimzamani-2018) on synthetic datasets revealed that (Zan-2022):

- FP performs well in the purely quantitative case with no conditioning but is not competitive in the mixed data case
- MS and RAVK are close to each other and have similar performance. Main drawback is that MS gives the value 0, or close to 0, to the estimator in some particular cases (as when the k-NN is always determined by Z)
- CMIh and LH behave well overall but LH is so slow that it can not be used when the dimension of Z exceeds 2
- The good behaviour of CMIh is further confirmed when used in conjunction with a local a permutation test (Zan-2022)

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- CMIh and LH behave well overall but LH is so slow that it can not be used when the dimension of Z exceeds 2
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→ *Discussion with Lei*



Illustration (2)

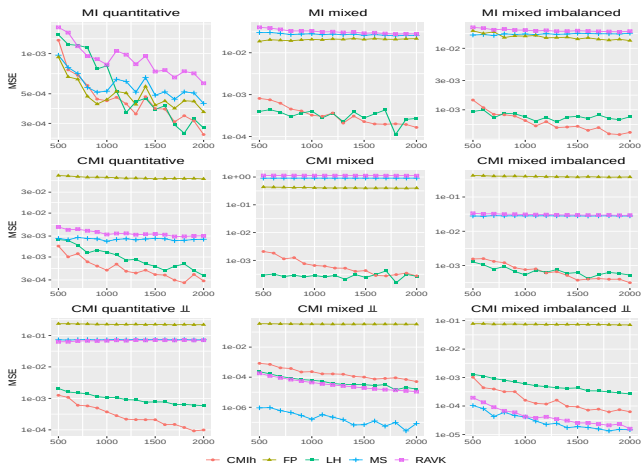


Figure: Synthetic data with known ground truth. MSE (on a log-scale) of each method with respect to the sample size (in abscissa)

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Conclusion

Exploring the use of (extended) summary graphs ...

- 1 Operational advantages: easier to infer, validate and come up with
- 2 Inferring such graphs from observational data
 - Independence estimation is a practical bottleneck (precise and fast)
- 3 Solving the identifiability problem in such graphs (on-going work)

for mixed time series

A few challenges

- Non stationarity
- Cyclic causal graphs (instantaneous relations)
- Efficient and effective methods with (some) theoretical guarantees with hidden confounders
- ...

Survey and Evaluation of Causal Discovery Methods for Time Series
(Assaad-2022(a); JAIR, IJCAI)

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- [Asaad-2021](#) K. Assaad, E. Devijver, E. Gaussier, A. Aït-Bachir. A Mixed Noise and Constraint-Based Approach to Causal Inference in Time Series. ECML/PKDD 2021
- [Asaad-2022\(a\)](#) C. K. Assaad, E. Devijver, E. Gaussier. Survey and evaluation of causal discovery methods for time series. Journal of Artificial Intelligence Research, 73: 767–819, feb 2022.
- [Asaad-2022\(b\)](#) C. K. Assaad, E. Devijver, E. Gaussier. Causal Discovery of Extended Summary Graphs in Time Series. UAI 2022
- [Asaad-2022\(c\)](#) C. K. Assaad, E. Devijver, E. Gaussier. Entropy-Based Discovery of Summary Causal Graphs in Time Series. Entropy 24(8), 2022
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